

# AN ALGORITHM FOR THE CHOICE OF OPTIMAL RESPONSE SURFACE DESIGNS

BY

M. SINGH, A. DEY AND R.K. MITRA

*I.A.S.R.I., New Delhi*

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## 1. INTRODUCTION AND PRELIMINARIES

Several optimality criteria in general regression problems have been discussed by Kiefer [2] and Kiefer and Wolfowitz [3]. Many designs are available in literature for fitting the second order response surfaces. In particular, central composite designs, Box-Behnken designs, uniform shell designs are recommended for practical purposes. Lucas [4] has tabulated the  $G$ -efficiencies and  $D$ -efficiencies of some designs of above types. The choice between two designs can be made on the basis of comparing the efficiencies of the two designs.

Lucas [4] made use of the computer for calculation of ultimate efficiencies of various designs which involves a long process. In the present paper we have evolved an algorithm for comparing various second order response surface designs.

Following Kiefer [2] a design can be treated as a measure  $\xi$  on  $\chi$  by  $\xi(x)$ , representing the proportion of observations taken at a point  $x$ ,  $\chi$  being the experimental region usually taken as  $[-1, 1]$ . In an  $N$  point design with  $n_i$  observations at  $x_i$  ( $\sum n_i = N$ ), we have

$$\xi(x_j) = \begin{cases} 0 & \text{if there are no observations at } x_j \\ n_j/N & \text{if } n_j > 0 \end{cases}$$

In a discrete  $N$  point design,  $\xi$  takes value  $1/N$  and defines an exact design.

Consider the usual regression set up given by

$$E(y_x) = \sum_{i=1}^p \beta_i f_i(x) = \underline{f}'(x) \underline{\beta}$$

$$V(y_x) = \sigma^2$$

where  $\underline{\beta} = (\beta_1, \dots, \beta_p)'$

is column vector of regression coefficients and  $\sigma^2$  is the per observation variance. Given an integer  $N$ , it is desired to select a set

$$\underline{x} = (x_1, \dots, x_N)'$$

of  $N$  points in  $\mathcal{X}$  for which the random variables

$$Y = (y_{x_1}, \dots, y_{x_N})'$$

are observed. We can also write

$$E(y_x) = \underline{f}'(x) \underline{\beta}$$

or

$$E(Y) = \underline{X} \underline{\beta}$$

where  $i$ -th row of  $\underline{X}$  is  $\underline{f}'(x_i)$ .

The estimated response at a point  $x \in \mathcal{X}$  along with its variance is given by

$$\hat{y}(x) = \underline{f}'(x) \hat{\beta}$$

$$V(\hat{y}(x)) = \sigma^2 \underline{f}'(x) (\underline{X}' \underline{X})^{-1} \underline{f}(x)$$

where  $\hat{\beta}$  is least square estimate of  $\beta$ .

Smith [5] proposed the criterion

$$\min_{\{x_i, i=1, 2, \dots, N\}^{x \in \mathcal{X}}} \max V(\hat{y}(x))$$

for optimal experimental design when considering the polynomial regression of degree  $p-1$  in one variable over the region  $\mathcal{X} = [-1, 1]$ . This criterion was later called as  $G$ -optimality by Kiefer and Wolfowitz [3]. They extended this criterion in general case to a design measure satisfying

$$\xi(x) \geq 0, x \in \mathcal{X}, \int_{\mathcal{X}} \xi(dx) = 1$$

Thus the variance function of estimated response at a point  $x$  is given by

$$d(x, \xi) = N \underline{f}'(x) (\underline{X}' \underline{X})^{-1} \underline{f}(x).$$

A design with measure  $\xi^*$  is said to be  $G$ -optimal if

$$\min_{\xi} \max_{x \in \mathcal{X}} d(x, \xi) = \max_{x \in \mathcal{X}} d(x, \xi^*).$$

A sufficient condition for satisfying above is

$$\max_{x \in \mathcal{X}} d(x, \xi^*) = p.$$

The  $G$ -efficiency of a design  $\xi$  is given in comparison to an  $G$ -optimal design by  $100 p / \max_{x \in \mathcal{X}} d(x, \xi)$ .

2. THE ALGORITHM

Suppose it is desired to fit the second order response surface given by

$$y = \beta_0 + \sum_{i=1}^v \beta_i x_i + \sum_{i=1}^v \beta_{ii} x_i^2 + \sum_{i < j} \beta_{ij} x_i x_j$$

under the experimental region  $X = \{x : x' x \leq 1\}$ .

In this case

$$p = (v+1)(v+2)/2.$$

We shall consider only the second order rotatable designs (SORD) introduced by Box and Hunter [1].

The expression of the variance of estimated response at point  $x_u = (x_{1u}, \dots, x_{vu})$  for the SORD  $\xi$  is

$$V(\hat{y}_u) = d(x, \xi) = \theta_0 + \theta_1 x'_u x_u + \theta_2 \left( x'_u x_u \right)^2$$

where  $\theta_0 = (v+2) \lambda_4 \sigma^2 / N \Delta$

$$\theta_1 = \{(1/\lambda_2) - (2\lambda_2/\Delta)\} \sigma^2 / N$$

$$\theta_2 = \{1 + (\lambda_2^2 - \lambda_4) / \Delta\} \sigma^2 / N$$

and

$$\Delta = (v+2) \lambda_4 - v \lambda_2^2$$

$$N \lambda_2 = \sum_u x_{iu}^2 \quad \text{for all } i=1, 2, \dots, v$$

$$N \lambda_4 = \sum_u x_{iu}^2 x_{ju}^2 \quad \text{for all } i < j=1, 2, \dots, v$$

$$\xi(x_u) = 1/N, \quad \text{for all } u=1, 2, \dots, N.$$

In order to obtain the factor combination at which the variance of estimated response is maximum, it is easily noticed that the variance expression is a function of  $x' x$ , the square of the distance of point  $x$  from origin. Thus

$$\max V(\hat{y}) = \begin{cases} \theta_0 & \text{if } \theta_1 + \theta_2 \leq 0 \\ \theta_0 + \theta_1 + \theta_2 & \text{if } \theta_1 + \theta_2 > 0 \end{cases}$$

that is according as the factor levels giving the maximum variance, are at the centre or at the surface of the hypersphere  $x' x = 1$ .

The condition  $\theta_1 + \theta_2 \leq 0$  can be expressed as

$$\lambda_4 \leq \lambda_4^0$$

$$\text{where } \lambda_4^0 = \left\{ \lambda_2^2 - (v+1) \lambda_2 / 2(v+2) \right\} / 2 \\ + \left[ \left\{ \lambda_2^2 - (v+1) \lambda_2 / 2(v+2) \right\}^2 + (v-1) \lambda_2^3 / 2(v+2) \right]^{1/2}$$

In the following, we shall obtain the condition for the choice of the best design. Suppose we have designs  $D_1$  and  $D_2$  in  $N_1$  and  $N_2$  points respectively. There are three different situations for comparisons.

(i) Both the  $\lambda_4$  parameters are less than their respective  $\lambda_4^0$  parameters. In this case, the design  $D_1$  is preferred to  $D_2$  if

$$\lambda_{42} / \lambda_{41} < \lambda_{22}^2 / \lambda_{21}^2$$

where  $\lambda_{4i}$  and  $\lambda_{2i}$  are  $\lambda_4$  and  $\lambda_2$  parameters of  $i$ th design  $D_i$ ,  $i=1, 2$ .

In a class of such designs  $\{D_j\}$  we shall choose  $D_i$  if

$$\min \left\{ \lambda_{2j}^2 / \lambda_{4j} \right\} = \lambda_{2i}^2 / \lambda_{4i}$$

If  $\lambda_{2j}=1$  for all  $j$ , the design which has the largest  $\lambda_4$  is preferred.

(ii) If  $\lambda_4$  parameter of  $D_1$  is less than  $\lambda_4^0$  for the design  $D_1$  and  $\lambda_4$  of  $D_2$  is greater than  $\lambda_4^0$  of  $D_2$ , then  $D_1$  is preferred to  $D_2$  if

$$N_1 \theta_{01} < N_2 (\theta_{02} + \theta_{12} + \theta_{22})$$

where  $\theta(\cdot)_i$ 's are the  $\theta(\cdot)$  parameter of  $D_i$ ,  $i=1, 2$ .

(iii) If  $\lambda_4$  parameters are greater than  $\lambda_4^0$  values for both the designs then  $D_1$  is preferred to  $D_2$  if

$$N_1 (\theta_{01} + \theta_{11} + \theta_{21}) < N_2 (\theta_{02} + \theta_{12} + \theta_{22})$$

Therefore in order to choose a best design among the set of available designs, one may check one of the above possibilities rather than making ultimate computation of  $G$ -efficiencies of all the designs as suggested in Lucas [4].

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## SUMMARY

An algorithm for the choice of most efficient (according to  $G$ -efficiency criterion) design among the set of a several competing designs has been worked out in case of second order rotatable designs.

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